

The cover time of two classes of random graphs

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1 Introduction

Let $G = (V, E)$ be a connected graph, let $|V| = n$, and $|E| = m$. A random walk \mathcal{W}_u , $u \in V$ on the undirected graph $G = (V, E)$ is a Markov chain $X_0 = u, X_1, \dots, X_t, \dots \in V$ associated to a particle that moves from vertex to vertex according to the following rule: the probability of a transition from vertex i , of degree d_i , to vertex j is $1/d_i$ if $\{i, j\} \in E$, and 0 otherwise. For $u \in V$ let C_u be the expected time taken for \mathcal{W}_u to visit every vertex of G . The cover time C_G of G is defined as $C_G = \max_{u \in V} C_u$. The cover time of connected graphs has been extensively studied. It is a classic result of Aleliunas, Karp, Lipton, Lovász and Rackoff [2] that $C_G \leq 2m(n-1)$. It was shown by Feige [11], [12], that for any connected graph G

$$(1 - o(1))n \ln n \leq C_G \leq (1 + o(1))\frac{4}{27}n^3.$$

The lower bound is achieved by (for example) the complete graph K_n , whose cover time is determined by the Coupon Collector problem.

In a previous paper [10] we studied the cover time of random graphs $G_{n,p}$ when $np = c \ln n$ where $c = O(1)$ and $(c-1) \ln n \rightarrow \infty$. This extended a result of Jonasson, who proved in [16] that when the expected average degree $(n-1)p$ grows faster than $\ln n$, **whp** a random graph has the same cover time (asymptotically) as the complete graph K_n , whereas, when

$np = \Theta(\ln n)$ this is not the case.

Theorem 1 [10] *Suppose that $np = c \ln n = \ln n + \omega$ where $\omega = (c-1) \ln n \rightarrow \infty$ and $c \geq 1$. If $G \in G_{n,p}$, then **whp**¹*

$$C_G \sim c \ln \left(\frac{c}{c-1} \right) n \ln n.$$

The notation $A_n \sim B_n$ means that $\lim_{n \rightarrow \infty} A_n/B_n = 1$.

We first consider random regular graphs:

Theorem 2 *Let $r \geq 3$ be constant. Let \mathcal{G}_r denote the set of r -regular graphs with vertex set $V = \{1, 2, \dots, n\}$. If G is chosen randomly from \mathcal{G}_r , then **whp***

$$C_G \sim \frac{r-1}{r-2} n \ln n.$$

Aldous [1] found the cover time of certain Cayley graphs. Once we have proved Theorem 2 we will see that some of Aldous's results can be obtained fairly easily. This connection will be discussed in the full paper.

We turn our attention to the preferential attachment graph $G_m(n)$ introduced by Barabási and Albert [4] as a simplified model of the WWW. The preferential attachment graph $G_m(n)$ is a random graph formed by adding a new vertex at each time step, with m edges which point to vertices selected at random with probability proportional to their degree. Thus at

¹A sequence of events \mathcal{E}_n occurs *with high probability whp* if $\lim_{n \rightarrow \infty} \Pr(\mathcal{E}_n) = 1$.

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time n there are n vertices and mn edges. We use the generative model of [7] (see also [8]) and build a graph sequentially as follows:

- At each time step t , we add a vertex v_t , and we add an edge from v_t to some other vertex u , where u is chosen at random according to the distribution:

$$\Pr(u = v_i) = \begin{cases} \frac{d_t(v_i)}{2t-1}, & \text{if } v_i \neq v_t; \\ \frac{1}{2t-1}, & \text{if } v_i = v_t; \end{cases}$$

where $d_t(v)$ denotes the degree of vertex v at time t .

- For some constant m , every m steps we contract the most recently added m vertices to form a single vertex.

Let $G_m(n)$ denote the random graph at time step mn after n contractions of size m . Thus $G_m(n)$ has n vertices and mn edges and may be a multi-graph.

We prove

Theorem 3 Whp *the preferential attachment graph $G = G_m(n)$ satisfies*

$$C_G \sim \frac{2m}{m-1} n \ln n.$$

The next section contains the heart of the proof of our Theorems. In it we establish a good estimate of the probability that the first visit of \mathcal{W} to a vertex v takes place at a time t . Once this is done, we can proceed to the proof of Theorem 2 in Section 3.

2 The first visit time lemma.

2.1 Convergence of the random walk

In this section G denotes a fixed connected graph with n vertices. u is some arbitrary vertex from which a walk \mathcal{W}_u is started. Let $\mathcal{W}_u(t)$ be the vertex reached at step t , let P be the matrix of transition probabilities of the walk and let $P_u^{(t)}(v) =$

$\Pr(\mathcal{W}_u(t) = v)$. Let π be the steady state distribution of the random walk \mathcal{W}_u .

2.2 Generating function formulation

Fix two

vertices u, v . Let h_t be the probability $\Pr(\mathcal{W}_u(t) = v) = P_u^{(t)}(v)$, that the walk \mathcal{W}_u visits v at step t . Let $H(s)$ generate h_t .

Similarly, considering the walk \mathcal{W}_v , starting at v , let r_t be the probability that this walk returns to v at step $t = 0, 1, \dots$. Let $R(s)$ generate r_t . We note that $r_0 = 1$.

Let $f_t(u \rightarrow v)$ be the probability that the first visit of the walk \mathcal{W}_u to v occurs at step t . If $u \neq v$ then $f_0(u \rightarrow v) = 0$. Let $F(s)$ generate $f_t(u \rightarrow v)$. Thus

$$(1) \quad H(s) = F(s)R(s).$$

Let T be the smallest positive integer such that for $t \geq T$,

$$(2) \quad \max_{x \in V} |P_u^{(t)}(x) - \pi_x| \leq n^{-3}$$

For $R(s)$ let

$$(3) \quad R_T(s) = \sum_{j=0}^{T-1} r_j s^j.$$

Thus $R_T(s)$ generates the probability of a return to v during steps $0, \dots, T-1$ of a walk starting at v . Similarly for $H(s)$, let

$$(4) \quad H_T(s) = \sum_{j=0}^{T-1} h_j s^j.$$

2.3 First visit time: Single vertex v

The following lemma should be viewed in the context that G is an n vertex graph which is part of a sequence of graphs with n growing to infinity. We prove it in greater generality than is needed for the proof of Theorem 2.

Lemma 4 Let T be as defined in (2) and

$$(5) \quad \lambda = \frac{1}{K_1 T}$$

for sufficiently large K_1 .

Suppose that for some constant $0 < \theta < 1$,

$$(a) \quad H_T(1) < \theta R_T(1).$$

$$(b) \quad \min_{|s| \leq 1 + \lambda} |R_T(s)| \geq \theta.$$

$$(c) \quad T\pi_v = o(1), T\pi_v = \Omega(n^{-2}).$$

Let

$$(6) \quad p_v = \frac{\pi_v}{R_T(1)(1 + O(T\pi_v))},$$

$$(7) \quad c_{u,v} = 1 - \frac{H_T(1)}{R_T(1)(1 + O(T\pi_v))},$$

where the values of the $1 + O(T\pi_v)$ terms are given implicitly in (14), (17) respectively. Then

$$(8) \quad f_t(u \rightarrow v) = c_{u,v} \frac{p_v}{(1 + p_v)^{t+1}} + O(e^{-\lambda t/2}) \quad \text{for all } t \geq T.$$

Proof Write

$$(9) \quad R(s) = R_T(s) + \widehat{R}_T(s) + \frac{\pi_v s^T}{1 - s},$$

where $R_T(s)$ is given by (3) and

$$\widehat{R}_T(s) = \sum_{t \geq T} (r_t - \pi_v) s^t$$

generates the error in using the stationary distribution π_v for r_t when $t \geq T$. Similarly, let

$$(10) \quad H(s) = H_T(s) + \widehat{H}_T(s) + \pi_v \frac{s^T}{1 - s}.$$

Note that for $Z = H, R$ and $|s| \leq 1 + o(1)$,

$$(11) \quad |\widehat{Z}(s)| = o(n^{-2}).$$

This is because the variation distance between the stationary and the t -step distribution decreases exponentially with t .

Using (9), (10) we rewrite $F(s) = H(s)/R(s)$ from (1) as $F(s) = B(s)/A(s)$ where

$$(12) \quad A(s) = \pi_v s^T + (1 - s)(R_T(s) + \widehat{R}_T(s)),$$

$$(13) \quad B(s) = \pi_v s^T + (1 - s)(H_T(s) + \widehat{H}_T(s)).$$

For real $s \geq 1$ and $Z = H, R$, we have

$$Z_T(1) \leq Z_T(s) \leq Z_T(1)s^T.$$

Let $s = 1 + \beta\pi_v$, where $\beta > 0$ is constant. Since $T\pi_v = o(1)$ we have

$$Z_T(s) = Z_T(1)(1 + O(T\pi_v)).$$

$T\pi_v = o(1)$ and $T\pi_v = \Omega(n^{-2})$ and $R_T(1) \geq 1 + r_2 > 1 + \frac{1}{n}$ implies that

$$A(s) = \pi_v(1 - \beta R_T(1)(1 + O(T\pi_v))).$$

It follows that $A(s)$ has a real zero at s_0 , where

$$(14) \quad s_0 = 1 + \frac{\pi_v}{R_T(1)(1 + O(T\pi_v))} = 1 + p_v,$$

say. We also see that

$$(15) \quad A'(s_0) = -R_T(1)(1 + O(T\pi_v)) \neq 0$$

and thus s_0 is a simple zero (see e.g. [9] p193). The value of $B(s)$ at s_0 is

$$(16) \quad B(s_0) = \pi_v \left(1 - \frac{H_T(1)}{R_T(1)(1 + O(T\pi_v))} + O(T\pi_v) \right) \neq 0.$$

Thus, from (6), (7)

$$(17) \quad \frac{B(s_0)}{A'(s_0)} = -p_v c_{u,v}.$$

Thus the (see e.g. [9] p195) the principal part of the Laurent expansion of $F(s)$ at s_0 is

$$(18) \quad f(s) = \frac{B(s_0)/A'(s_0)}{s - s_0}.$$

To approximate the coefficients of the generating function $F(s)$, we now use a standard technique for the asymptotic expansion of power series (see e.g. [19] Th 5.2.1).

We prove below that $F(s) = f(s) + g(s)$, where $g(s)$ is analytic in C_λ and that $M = \max_{s \in C_\lambda} |g(s)| = O(1)$.

Let $a_t = [s^t]g(s)$, then (see e.g. [9] p143), $a_t = g^{(t)}(0)/t!$. By the Cauchy Inequality (see e.g. [9] p130) we have that $|g^{(t)}(0)| \leq Mt!/(1 + \lambda)^t$ and thus

$$|a_t| \leq \frac{M}{(1 + \lambda)^t} = O(e^{-t\lambda/2}).$$

As $[s^t]F(s) = [s^t]f(s) + [s^t]g(s)$ and $[s^t]1/(s - s_0) = -1/(s_0)^{t+1}$ we have

$$(19) \quad [s^t]F(s) = \frac{-B(s_0)/A'(s_0)}{s_0^{t+1}} + O(e^{-t\lambda/2}).$$

Thus, we obtain

$$[s^t]F(s) = c_{u,v} \frac{p_v}{(1 + p_v)^{t+1}} + O(e^{-t\lambda/2}),$$

which completes the proof of (8).

We now prove that s_0 is the only zero of $A(s)$ inside the circle C_λ . We use Rouché's Theorem (see e.g. [9]), the statement of which is as follows: *Let two functions $\phi(z)$ and $\gamma(z)$ be analytic inside and on a simple closed contour C . Suppose that $|\phi(z)| > |\gamma(z)|$ at each point of C , then $\phi(z)$ and $\phi(z) + \gamma(z)$ have the same number of zeroes, counting multiplicities, inside C .*

Let the functions $\phi(s), \gamma(s)$ be given by $\phi(s) = (1 - s)R_T(s)$ and $\gamma(s) = \pi_v s^T + (1 - s)\widehat{R}_T(s)$.

$$|\gamma(s)|/|\phi(s)| \leq \frac{\pi_v(1 + \lambda)^T}{\theta} + \frac{|\widehat{R}_T(s)|}{\theta} = o(1).$$

As $f(s) + g(s) = A(s)$ we conclude that $A(s)$ has only one zero inside the circle C_λ . This is the simple zero at s_0 . \square

Corollary 5 *Let $\mathbf{A}_t(v)$ be the event that \mathcal{W}_u has not visited v by step t . Then for $t \geq T$,*

$$\Pr(\mathbf{A}_t(v)) = \frac{c_{u,v}}{(1 + p_v)^t} + O(e^{-\lambda t/2}).$$

Proof We use Lemma 4 and

$$\Pr(\mathbf{A}_t(v)) = \sum_{\tau > t} f_\tau(u \rightarrow v).$$

3 Random regular graphs

From here on, we replace $R_T(1), H_T(1)$ by the notation R_v, H_v .

We start with some typical properties of a random regular graph. Let

$$\omega = \lfloor \ln \ln \ln n \rfloor.$$

Say a cycle C is *small* if $|C| \leq 2\omega + 1$. An r -regular graph G is *nice* if

P1. G is connected.

P2. The second eigenvalue of the adjacency matrix of G is at most $2\sqrt{r-1} + \epsilon$, where $\epsilon > 0$ is arbitrarily small ($\epsilon = 1/10$ is small enough). This implies that $T = O(\ln n)$.

P3. There are at most $r^{2\omega}$ vertices on small cycles.

P4. No pair of small cycles are within distance 3ω of each other.

In the full paper we show

Theorem 6 *Let $r \geq 3$ be a constant and let G be chosen uniformly from the set \mathcal{G}_r r -regular graphs with vertex set $[n]$. Then G is nice **whp**.*

Assume from now on that G is a nice regular graph.

For $v \in V$ let H_v be the sub-graph induced by the vertices at distance 2ω or less from v .

Definition 1 *We say v is locally tree-like if H_v is a tree.*

Lemma 7 *If v is locally tree-like then*

$$R_v = \frac{r-1}{r-2} + o(\omega^{-1}).$$

Proof Let T_r be the infinite r -regular tree, rooted at v . Let \mathcal{X} be a random walk on T_r starting at v . Let ρ_i be the probability that \mathcal{X} is at v at

step i . Now we can project the walk \mathcal{X} onto a walk on $\{0, 1, 2, \dots\}$ where the particle moves right with probability $q = \frac{r-1}{r}$ and left with probability $p = \frac{1}{r}$. Let E_i be the expected number of visits to 0 for such a walk starting at i . Then

$$E_0 = 1 + E_1 = 1 + E_0 p / q.$$

This is because E_1 is E_0 times the expected number of visits to 0 between right moves from 1. Solving gives

$$(20) \quad \sum_{i=0}^{\infty} \rho_i = E_0 = \frac{r-1}{r-2}.$$

Note next that $\rho_{2i+1} = 0, \rho_{2i} \leq \gamma_i = \binom{2i}{i} \left(\frac{r-1}{r^2}\right)^i$ bounds the number of walks that place the particle at 0 at time $2i$. Therefore,

$$(21) \quad \sum_{i=\omega+1}^{\infty} \rho_i \leq \sum_{j=\omega/2}^{\infty} \gamma_j = o(\omega^{-1}).$$

We compare this with R_v . First observe that $r_i = \rho_i$ for $i \leq \omega$. Then from e.g. [15], we see that

$$(22) \quad \sum_{i=\omega+1}^T r_i \leq \sum_{i=\omega+1}^T (\pi_v + \lambda_{\max}^i) = o(\omega^{-1}).$$

The lemma now follows from (20) and (21). \square

Lemma 8 *If v is locally tree-like then $|R_T(s)| \geq 1/(2(1+e))$ for $|s| \leq 1 + \lambda$*

Proof $R_T(s)$ generates the expected number of returns to v in T . Assuming v is locally tree-like

$$R_T(s) = \frac{1}{1-F(s)} + Q(s)$$

where $F(s)$ generates the first return probability in the tree T_r and $Q(s)$ is a correction. Thus

$$|R_T(s)| \geq \frac{1}{|1-F(s)|} - |Q(s)|.$$

and

$$|1-F(s)| \leq 1 + |F(s)| \leq 1 + (1+\lambda)^T,$$

as $|F(s)| \leq \sum_{j \geq 0} \gamma_j |s|^{2j} < 1$.

Write $Q(s) = Q_1(s) - Q_2(s)$ where

$$Q_1(s) = q_{\omega+1} s^{\omega+1} + \dots + q_{T-1} s^{T-1}$$

corrects returns due to non tree-like structure of G at steps $\omega+1, \dots, T-1$. $Q_2(s) = r_T s^T + \dots + r_t s^t + \dots$ is the tail of $M(s) = 1/(1-F(s))$ above $T-1$.

$\gamma_j \leq \left(\frac{4(r-1)}{r^2}\right)^j$ implies that the radius of convergence of $M(s)$ is $d > 1$, $r_t = O(d^{-t})$. It was proved in (22) that $Q_1(1) = O(1/\omega)$ so that $|Q_1(s)| \leq (1+\lambda)^T O(1/\omega)$ and for $1+\lambda < d$

$$|Q_2(s)| = O\left(\frac{((1+\lambda)/d)^T}{1-(1+\lambda)/d}\right) = O\left(\frac{1+\lambda}{d}\right)^T.$$

Thus, with $\alpha = (1+\lambda)^T$

$$\begin{aligned} |R_T(s)| &\geq \frac{1}{|1-F(s)|} - |Q(s)| \\ &\geq \frac{1}{1+\alpha} - \alpha O\left(\frac{1}{\omega}\right), \end{aligned}$$

Thus for $\lambda = 1/T$ we have $|R_T(s)| \geq 1/(2(1+e))$ for $|s| \leq 1 + \lambda$. \square

Finally we note:

Lemma 9 *For nice graphs, $\frac{H_v}{R_v} \leq \frac{9}{10}$.*

Proof Let f'_t be the probability that \mathcal{W}_u has a first visit to v at time t . As $H(s) = F(s)R(s)$ we have

$$\begin{aligned} H_T(1) &\leq \Pr(\mathcal{W}_u \text{ visits } v \text{ by time } T-1) R_T(1) \\ &= R_T(1) \sum_{t=1}^{T-1} f'_t. \end{aligned}$$

Now if $\tau_0 = \lfloor 2 \ln \lambda_{\max}^{-1} \ln \ln n \rfloor$ then

$$\sum_{t=\tau_0}^{T-1} f'_t \leq \sum_{t=\tau_0}^{T-1} (\pi_v + \lambda_{\max}^t) = o(1).$$

We now estimate $\sum_{t=0}^{\tau_0} f'_t$, the probability that \mathcal{W}_u visits v by time τ_0 . Let v_1, v_2, \dots, v_r be the neighbours of v and let w be the first neighbour of v visited

by \mathcal{W}_u . Then

$$\begin{aligned} \Pr(\mathcal{W}_u \text{ visits } v \text{ by time } \tau_0) &= \\ \sum_{i=1}^r \Pr(\mathcal{W}_u \text{ visits } v \text{ by time } \tau_0 \mid w = v_i) \Pr(w = v_i) &\leq \\ \sum_{i=1}^r \Pr(\mathcal{W}_{v_i} \text{ visits } v \text{ by the time } \tau_0) \Pr(w = v_i). \end{aligned}$$

So it suffices to prove the lemma when u is a neighbour of v . If $G_l(u)$ is a tree then we can argue as in Lemma 7. Let ψ be the probability that a particle at the root of T_r ever returns to the root. The expected number of visits is

$$\frac{r-1}{r-2} = \sum_{k=1}^{\infty} k \psi^{k-1} (1-\psi) = \frac{1}{1-\psi}.$$

So $\psi = \frac{1}{r-1}$ and

$$\begin{aligned} \Pr(\mathcal{W}_u \text{ does not visit } v \text{ by time } \tau_0) &\geq \\ \frac{r-1}{r} (1-\psi - o(1)) &= \frac{r-2}{r} - o(1). \end{aligned}$$

If $G_l(u)$ contains a cycle C then let $e = (\xi, \eta)$ be an edge of C not incident with u and let T_u be the tree $G_l(u) - e$. Let $N'(u) = \{u_1, u_2, \dots, u_s\}$, $s \in \{r-2, r-1\}$ be the neighbours of u which are not on a shortest path from ξ or η to u in T_u . $|N'(u) \setminus \{v\}| \geq r-3$ and so

$$\begin{aligned} \Pr(\mathcal{W}_u \text{ does not visit } v \text{ by time } \tau_0) &\geq \\ \frac{r-3}{r} (1-\psi - o(1)) &= \frac{(r-2)(r-3)}{r(r-1)} - o(1). \end{aligned}$$

This leaves the case $r=3$ and $N'(u) = \{v\}$. With probability $\frac{2}{3}$ we have $\mathcal{W}_u(1) \neq v$. If ξ or η is reached (possibly $N(u) = \{v, \xi, \eta\}$), then with probability $\frac{1}{3}$ the next move is away from u and $1-\psi - o(1)$ bounds the probability that there is no return to ξ or η . Hence

$$\Pr(\mathcal{W}_u \text{ does not visit } v \text{ by time } \tau_0) \geq \frac{2}{9} (1-\psi - o(1))$$

completing the proof of the lemma. \square

4 Cover time of nice graphs

We now prove that

$$C_G \sim \frac{r-1}{r-2} n \ln n.$$

Assume that $u, v \in V$ and that v is tree-like. Section 3 establishes that the conditions of Lemma 4 hold, and gives values for the parameters c_{uv}, p_v given by (6), (7).

Hence, the probability that \mathcal{W}_u has not visited v by some step $t \geq T$ (see Corollary 5) is given by

$$\Pr(\mathbf{A}_t(v)) = (1 + o(1)) c_{uv} e^{-tp_v} + O(\lambda^{-1} e^{-\lambda t/2}).$$

Here $c_{uv} = \Theta(1)$, $\lambda = \Theta(1/\ln n)$ and

$$p_v = \frac{r-2}{(r-1)n} (1 + o(\omega^{-1})).$$

4.1 Upper bound on cover time

Let $t_0 = \lceil (1 + \sigma^{-1}) \frac{r-1}{r-2} n \ln n \rceil$. We prove that for nice graphs, for any vertex $u \in V$,

$$(23) \quad C_u \leq t_0 + o(t_0).$$

Let $T_G(u)$ be the time taken to visit every vertex of G by the random walk \mathcal{W}_u . Let U_t be the number of vertices of G which have not been visited by \mathcal{W}_u at step t . We note the following:

$$(24) C_u = \mathbf{E} T_G(u) = \sum_{t>0} \Pr(T_G(u) \geq t),$$

$$(25) \Pr(T_G(u) > t) = \Pr(U_t > 0) \leq \min\{1, \mathbf{E} U_t\}.$$

It follows from (24), (25) that for all t

$$(26) \quad C_u \leq t + \sum_{s \geq t} \mathbf{E} U_s = t + \sum_{v \in V} \sum_{s \geq t} \Pr(\mathbf{A}_s(v)).$$

Let V_1 be the set of locally tree-like vertices and let $V_2 = V - V_1$. If G is nice then $|V_2| \leq r^{3\omega}$ for there are at most r^ω vertices within distance ω of a particular

vertex in a small cycle, and at most $r^{2\omega}$ vertices on small cycles.

For $v \in V_1$ we have

$$\begin{aligned} & \sum_{s \geq t_0} \Pr(\mathbf{A}_s(v)) \\ & \leq (1 + o(1))e^{-t_0 p_v} \sum_{s \geq t_0} e^{-(s-t_0)p_v} + O(\lambda^{-2}e^{-\lambda t_0/2}) \\ & \leq \pi_v^{-1} e^{-t_0 p_v} \\ & \leq 3 \frac{r-1}{r-2}. \end{aligned}$$

Furthermore, we see that in particular,

$$(27) \quad \Pr(\mathbf{A}_{5n}(v)) \leq 2e^{-1}.$$

Suppose next that $v \in V_2$. We can find $w \in V_1$ such that $\text{dist}(v, w) \leq \omega$. So from (27), with $\nu = 5n + \omega$, we have

$$\Pr(\mathbf{A}_\nu(v)) \leq 1 - (1 - 2e^{-1})r^{-\omega}$$

since if our walk visits w , it will with probability at least $r^{-\omega}$ visit v within the next ω steps. Thus if $\gamma = (1 - 2e^{-1})r^{-\omega}$,

$$\begin{aligned} \sum_{s \geq t_0} \Pr(\mathbf{A}_s(v)) & \leq \sum_{s \geq t_0} (1-\gamma)^{\lfloor s/\nu \rfloor} \leq \sum_{s \geq t_0} (1-\gamma)^{s/(2\nu)} \\ & = \frac{(1-\gamma)^{t_0/(2\nu)}}{1 - (1-\gamma)^{1/(2\nu)}} \leq 3\nu\gamma^{-1}. \end{aligned}$$

Thus, for all $u \in V$,

$$\begin{aligned} C_u & \leq t_0 + 3 \frac{r-1}{r-2} |V_1| + 3|V_2| \nu \gamma^{-1} \\ & = t_0 + O(r^{4\omega} n) = t_0 + o(t_0). \end{aligned}$$

4.2 Lower bound on cover time

For any vertex u , we can find a set of vertices S such that at time $t_1 = t_0(1 - \epsilon)$, $\epsilon \rightarrow 0$, the probability the set S is covered by the walk \mathcal{W}_u tends to zero. Hence $T_G(u) > t_1$ **whp** which implies that $C_G \geq t_0 - o(t_0)$.

We construct S as follows. Let $S \subseteq V_1$ be some maximal set of locally tree-like vertices all of which

are at least distance $2\omega + 1$ apart. Thus $|S| \geq (n - r^{3\omega})r^{-(2\omega+1)}$.

Let $S(t)$ denote the subset of S which has not been visited by \mathcal{W}_u after step t . Now, provided $t \geq T$

$$\mathbf{E} |S(t)| \geq (1 - o(1)) \sum_{v \in S} \left(\frac{c_{u,v}}{(1 + p_v)^t} + o(n^{-2}) \right).$$

Let u be a fixed vertex of S . Let $v \in S$ and let H_v be given by (4), then (32) implies that

$$(28) \quad H_v \leq \sum_{t=\omega}^{T-1} (\pi_v + \lambda_{\max}^t) = o(1).$$

Thus $c_{uv} = 1 - o(1)$. Setting $t = t_1 = (1 - \epsilon)t_0$ where $\epsilon = 2\omega^{-1}$, we have

$$(29) \quad \begin{aligned} \mathbf{E} |S(t_1)| & = (1 + o(1)) |S| e^{-(1-\epsilon)t_0 p_v} \\ & \geq n^{1/\omega}. \end{aligned}$$

Let $Y_{v,t}$ be the indicator for the event that \mathcal{W}_u has not visited vertex v at time t . Let $Z = \{v, w\} \subset S$. We can show (proof omitted) that that for $v, w \in S$

$$(30) \quad \mathbf{E} (Y_{v,t_1} Y_{w,t_1}) = \frac{c_{u,Z}}{(1 + p_Z)^{t+2}} + o(n^{-2}),$$

where $c_{u,Z} \sim 1$ and $p_Z \sim 2(r-2)/(n(r-1))$. Thus

$$(31) \quad \mathbf{E} (Y_{v,t_1} Y_{w,t_1}) = (1 + o(1)) \mathbf{E} (Y_{v,t_1}) \mathbf{E} (Y_{w,t_1}).$$

It follows from (29) and (31), that

$$\begin{aligned} \Pr(S(t_1) \neq \emptyset) & \geq \frac{(\mathbf{E} |S(t_1)|)^2}{\mathbf{E} |S(t_1)|^2} \\ & = \frac{1}{\frac{\mathbf{E} |S_{t_1}| (|S_{t_1}| - 1)}{(\mathbf{E} |S(t_1)|)^2} + (\mathbf{E} |S_{t_1}|)^{-1}} = 1 - o(1). \end{aligned}$$

□

5 Preferential Attachment Graph

5.1 The random graph $G_m(n)$

Lemma 10 (a) *Suppose that $0 < \alpha < \beta < 2/3$.*

Then

$$\Pr(\exists i \leq n^\alpha : d(i) \leq n^{1/2-\alpha/2-\beta}) = o(1).$$

(b)

$$\Pr(\exists s, t : d_t(s) \geq (t/s)^{1/2}(\ln n)^3) = O(n^{-3}).$$

□

Suppose that v is locally tree-like. We say that v is *locally regular* if H_v is a tree of depth 2ω , rooted at v , in which every non-leaf has branching factor m .

Lemma 11 Whp, $G_m(n)$ contains at least $n^{1-o(1)}$ locally regular vertices.

A small cycle is *light* if it contains no vertex $v \leq n^{1/10}$ (it has no “heavy” vertices), otherwise it is *heavy*.

Lemma 12 Whp $G_m(n)$ does not contain a small cycle within 10ω of a light cycle.

Lemma 13 Whp $G_m(n)$ does not contain a vertex $v \geq n^{3/5}$ which is within distance 10ω of 2 distinct small cycles.

We also need to deal with the possibility that $G_m(n)$ contains many cycles.

Lemma 14 Whp $G_m(n)$ contains at most $(\ln n)^{5\omega}$ small cycles.

Lemma 15 Whp there are at most $O(n^{1/2+o(1)})$ non tree-like vertices.

The *conductance* Φ of the walk \mathcal{W}_u is defined by

$$\Phi = \min_{\pi(S) \leq 1/2} \frac{e(S; \bar{S})}{d(S)}.$$

Mihail, Papadimitriou and Saberi [18] proved that the *conductance* Φ of the walks \mathcal{W} are bounded below by some absolute constant. Now it follows from Jerrum and Sinclair [15] that

$$(32) \quad |\tilde{P}_u^{(t)}(x) - \pi_x| \leq (\pi_x/\pi_u)^{1/2}(1 - \Phi^2/2)^t.$$

For sufficiently large K , the RHS above will be $O(n^{-10})$ at τ_0 . We remark that there is a technical point here. The result of [15] assumes that the walk is *lazy*, and only makes a move to a neighbour with probability $1/2$ at any step. This halves the conductance but we still have $T = O(\ln n)$ in (2). It doubles the covertime. It also asymptotically doubles the values R_v . Otherwise, it has a negligible effect on the analysis and we will ignore this for the rest of the paper and continue as though there are no lazy steps.

6 Cover time of $G_m(n)$

6.1 Parameters

Assume now that $G_m(n)$ (i) has $n^{1-o(1)}$ locally regular vertices, (ii) $d(s) \geq n^{1/4}$ for $s \leq n^{1/10}$, (iii) $d(s) \geq n^{1/25}$ for $s \leq n^{3/5}$, (iv) no small cycle close to a light cycle, (v) no $v \geq n^{3/5}$ within distance 10ω of 2 distinct small cycles, (vi) $O((\ln n)^{5\omega})$ small cycles and (vii) $O(n^{1/2+o(1)})$ non tree-like vertices.

Lemma 16 Suppose that v is locally-tree-like. Then

$$(a) \quad R_v \leq \frac{d(v)}{m-1}.$$

$$(b) \quad d(v) \geq m + 1 \quad \text{implies} \quad R_v \leq \frac{d(v)(m+m^{-1}-1)}{d(v)(m+m^{-1}-2)+m^{-1}-1}$$

Proof We first define an infinite tree T_v^* by taking the tree T'_v defined by the first $\omega + 1$ levels of H_v and then rooting a copy of the infinite tree T_m^∞ which has branching factor m from each leaf of T'_v . Thus if v is locally regular, T_v itself is an infinite tree with branching factor m , rooted at v .

Let R_v^* be the expected number of visits to v for infinite random walks \mathcal{W}_v^* on T_v^* , started at v , making null moves with probability γ and making no null moves respectively. We argue first that

$$(33) \quad |R_v - R_v^*| = o(\omega^{-1}).$$

Let $r_t^* = \mathbf{Pr}(\mathcal{W}_v^*(t) = v)$. Then

$$\begin{aligned}
|R_v - R_v^*| &\leq \sum_{t=\omega+1}^T r_t + \sum_{t=\omega+1}^{\infty} r_t^* \\
(34) \quad &\leq o(\omega^{-1}) + \sum_{t=\omega/10}^{\infty} \mathbf{1}_{t \text{ even}} \binom{t}{t/2} \frac{(m^2 - m)^{t/2}}{(m^2 - m + 1)^t} \\
&= o(\omega^{-1}).
\end{aligned}$$

Explanation of (34): The first term follows directly from (32): $\sum_{t=\omega+1}^T r_t \leq T\pi_v + \sum_{t=\omega+1}^T (1 - \Phi^2/2)^t$. We bound the second sum by considering a walk \mathcal{Y} on $\{0, 1, 2, \dots\}$ which at each time step moves right with probability $q_{\text{odd}} = \frac{m-1}{m}$ when at odd values, $q_{\text{even}} = \frac{m}{m+1}$ when at even values and moves left with probability $p_{\text{parity}} = 1 - q_{\text{parity}}$, $\text{parity}=\text{odd, even}$, except at the origin, when it always moves right. We couple $\mathcal{Y}(t)$ with the distance of $\mathcal{W}_v^*(t)$ from v . Our choice for $p_{\text{odd}}, p_{\text{even}}$ is determined by the fact that if a vertex w in the first $\omega+1$ levels of H_v has branching factor $m-1$ then its ancestors have branching factor $\geq m$. We maximise R_v by keeping branching factors small and so the largest R_v is achieved by having branching factors $m-1, m$ alternating on any path from v . This leads to \mathcal{Y} .

Thus r_t^* is at most the probability that $\mathcal{Y}(t) = 0$. We now bound this latter probability. We observe that it is bounded by the probability that another walk \mathcal{Y}_1 is at the origin after t steps. Here \mathcal{Y}_1 is the walk on $\{0, \pm 1, \pm 2, \dots\}$ where the particle moves right with probabilities $q_{\text{odd}}, q_{\text{even}}$ and left with probabilities $p_{\text{odd}}, p_{\text{even}}$ i.e. there is no barrier at the origin. We can couple $\mathcal{Y}, \mathcal{Y}_1$ so that $\mathcal{Y}(t) \geq |\mathcal{Y}_1(t)|$. When $\mathcal{Y}_1(t) > 0$ we can move them in the same direction and when $\mathcal{Y}_1 < 0$ then we can move \mathcal{Y} further from the origin whenever \mathcal{Y}_1 moves further from the origin.

Finally, consider the walk \mathcal{Y}_2 that \mathcal{Y}_1 induces on the even integers. The non-trivial moves are right with probability $\frac{m^2-m}{m^2-m+1}$ and left with probability $\frac{1}{m^2-m+1}$. The probability that t non-trivial moves yields a return is precisely $\mathbf{1}_{t \text{ even}} \binom{t}{t/2} \frac{(m^2-m)^{t/2}}{(m^2-m+1)^t}$. Now the probability that there are at most $\omega/10$ non-

trivial moves is exponentially small, in ω and this can be absorbed into the $o(\omega^{-1})$ term.

Let $b_w, w \in T_v^*$ be the branching factor at w i.e. $b_v = d_v$ and $b_w = d_w - 1$ if w is not the root. Further, if w is in the first ω levels let $b_w = b_w^+ + b_w^-$ where b_w^+ is the number of descendants w' of w with $w > w'$ i.e. w chose w' in the construction of $G_m(n)$. If w is at a higher level, we take $b_w = b_w^+ = m$ and $b_w^- = 0$.

Let \widehat{T}_w be the sub-tree of T_v^* rooted at vertex w . (Thus $\widehat{T}_v = T_v^*$). Let ρ_w denote the probability that a random walk on \widehat{T}_w which starts at w ever returns to w . Our aim is to estimate ρ_v and use

$$(35) \quad R_v^* = \frac{1}{1 - \rho_v}.$$

Let $C(w)$ denote the children of w in T_v^* . We use the following recurrence: The parameter k counts the number of returns to x .

$$\begin{aligned}
\rho_w &= \\
&1 - \frac{1}{b_w} \sum_{x \in C(w)} \sum_{k \geq 0} \left(1 - \frac{1}{d_x}\right) \left(\rho_x \left(1 - \frac{1}{d_x}\right)\right)^k (1 - \rho_x) \\
&= 1 - \frac{1}{b_w} \sum_{x \in C(w)} \frac{\left(1 - \frac{1}{d_x}\right) (1 - \rho_x)}{1 - \rho_x \left(1 - \frac{1}{d_x}\right)} \\
&= \frac{1}{b_w} \sum_{x \in C(w)} \frac{1}{b_x + 1 - \rho_x b_x}.
\end{aligned}$$

We see immediately that if T_v^* is a regular tree with branching factor $m \geq 2$ then, with $\rho_w = \rho$ for all w ,

$$\rho = \frac{1}{m+1-\rho m} \text{ and hence } \rho = \frac{1}{m}$$

and this deals with the locally regular case.

Now define b_w^+ to be the number of children x of w with $x < w$. These are the children chosen by w . Let $b_w^- = b_w - b_w^+$.

We will now prove the following by induction on $\omega + 1 - \ell_w$, where $\ell_w \leq \omega + 1$ is the level of w in the tree.:

(a) $b_w = m - 1$ implies $\rho_w \leq \frac{1}{m}$.

(b) $b_w \geq m + 1$ implies $\rho_w \leq \frac{1}{b_w} \left(1 + \frac{b_w - m}{m + m^{-1} - 1}\right)$.

(c) $b_w = b_w^+ = m$ implies $\rho_w \leq \frac{1}{m}$.

(d) $b_w = b_w^+ + 1 = m$ implies

$$\rho_w \leq \frac{1}{m} \left(\frac{m-1}{m} + \frac{m}{m^2 - m + 1}\right)$$

The base case will be $\ell_w = \omega + 1$. For which, Case (c) applies and the induction hypothesis holds from the locally regular case.

The lemma follows from this since only cases (b),(c),(d) can apply to the root v , in which case $b_v = d(v)$.

Let us now go through the inductive step. Let us assume these conditions apply to $x \in C(w)$ and then we find that in these cases:

(a) $b_x + 1 - b_x \rho_x \geq m + \frac{1}{m} - 1$.

(b) $b_x + 1 - b_x \rho_x \geq m + (b_x - m) \left(1 - \frac{1}{m + m^{-1} - 1}\right) \geq m$.

(c) $b_x + 1 - b_x \rho_x \geq m$.

(d) $b_x + 1 - b_x \rho_x \geq m + \frac{1}{m} - \frac{m}{m^2 - m + 1}$.

Case (a): In this case $b_w = b_w^+$ and only cases (b),(c) are possible for $x \in C(w)$. In which case $b_x + 1 - b_x \rho_x \geq m$ for $x \in C(w)$.

Case (b): In $C(w)$ we have b_w^+ cases of (b) or (c) and b_w^- cases of (a),(b) or (d). In the first case we have $b_x + 1 - b_x \rho_x \geq m$. In the second case we have $b_x + 1 - b_x \rho_x \geq m + m^{-1} - 1$. Thus

$$\rho_w \leq \frac{1}{b_w} \left(\frac{b_w^+}{m} + \frac{b_w^-}{m + m^{-1} - 1}\right)$$

Sub-case (i): $b_w^+ = m$.

$$\rho_w \leq \frac{1}{b_w} \left(1 + \frac{b_w - m}{m + m^{-1} - 1}\right).$$

Sub-case (ii): $b_w^+ = m - 1$.

$$\rho_w \leq \frac{1}{b_w} \left(1 - \frac{1}{m} + \frac{b_w - m + 1}{m + m^{-1} - 1}\right).$$

Case (c): This follows as in Case (a).

Case (d): In $C(w)$ we have $m - 1$ cases of (b) or (c) and one case of (a),(b) or (d). Thus

$$\rho_w \leq \frac{1}{m} \left(\frac{m-1}{m} + \frac{1}{m + m^{-1} - 1}\right)$$

as is to be shown. \square

Lemma 17 *Suppose that either*

(i) G_v contains a unique light cycle C_v , that $v \notin C_v$ and that the shortest path $P = (w_0 = v, w_1, \dots, w_k)$ from v to C_v is such that $\max\{d(w_1), \dots, d(w_k)\} \geq \omega^3$, or

(ii) that H_v contains only heavy cycles. Then

(a) $R_v \leq \frac{d(v)}{m-1}$.

(b) $d(v) \geq m + 1$ implies $R_v \leq \frac{d(v)(m+m^{-1}-1)}{d(v)(m+m^{-1}-2)+m^{-1}-1}$

Proof

(a) Let w be the first vertex on the path from v to C_v which has degree at least ω^3 . Let H'_v be obtained from H_v by deleting those vertices, other than w , whose only path to v in H_v goes through w . Let R'_v be the expected number of returns to v in a random walk of length ω on H'_v where w is an absorbing state. We claim that

$$(36) \quad R_v \leq R'_v + o(\omega^{-1}).$$

Once we verify this, the proof of (a) follows from the proof of Lemma 16 i.e. embed the tree H'_v in an infinite tree by rooting a copy of T_m^∞ at each leaf. To verify (36) we couple random walks on H_v, H'_v until w is visited. In the latter the process stops. In the former, we find that when at w , the probability we get closer to v in the next step is at most ω^{-3} and so the expected number of returns from now on is at most $\omega \times \omega^{-3}$ and (36) follows.

(b) Now consider the case where H_v contains only heavy cycles. We argue first that a random walk of length ω that starts at v might as well terminate if it reaches a vertex $w \leq n^{1/10}$, $w \neq v$. We can assume

$d(w) \geq n^{1/4}$. Now we can assume from Lemma 14 at least $n_0 = n^{1/4} - (\ln n)^{5\omega}$ of the edges incident with w are not in cycles contained in H_v . But then a walk that arrives at w has a more than $\frac{n_0}{n^{1/4}}$ chance of entering a sub-tree T_w of H_v rooted at w for which every vertex is separated from v by w . But then the probability of leaving T_w in ω steps is $O(\omega(\ln n)^{5\omega}/n^{1/4})$ and so once a walk has reached w , the expected number of further returns to v is $o(\omega^{-1})$. We can therefore remove T_w from H_v and then replace an edge (x, w) by an edge (x, w_x) and make all the vertices w_x absorbing. Repeating this argument, we are left with a tree to which we can apply the argument of Lemma 16. \square

Note that if $v \in V_B$ then no bound on R_v has been established:

$$V_B = \{v : G_v \text{ contains a unique light cycle } C_v \text{ and the path from } v \text{ to } C_v \text{ contains no vertex of degree at least } \omega^3\}$$

However, for these it suffices to prove

Lemma 18 *If $v \in V_B$ then $R_v \leq (\ln n)^{2/3}$.*

We will also need to show the following:

Lemma 19 *If $v \geq n^{3/5}$ then*

- (a) $|R_T(s)| \geq 1/(2(1+e))$ for $|s| \leq 1 + \lambda$.
- (b) $H_v < C_m R_v + o(1)$ where $C_m < 1$.

One of the problems in proving this lemma arises from the existence of non-locally-tree-like vertices. This problem is ameliorated by restricting attention to $v \geq n^{3/5}$ and using Lemma 13. For $v < n^{3/5}$, we know $d(v) \geq n^{1/25}$ and so after a walk of length $(\ln n)^3$ there is an $\Omega(n^{-1/25})$ chance of being at v . Thus v will be visited in $O(n^{24/25}(\ln n)^5)$ time **qs**.

6.2 Upper bound on cover time

Let $t_0 = \lceil \frac{2m}{m-1} n \ln n \rceil$. We prove that **whp**, for $G_m(n)$, for any vertex $u \in V$, $C_u \leq t_0 + o(t_0)$.

Arguing as in (26)

$$(37) \quad C_u \leq t + o(1) + \sum_{v \in V_B} \sum_{s \geq t} \Pr(\mathbf{A}_s(v)) + (1 + O(T\pi_v)) \times \sum_{\substack{v \in V \setminus V_B \\ v \geq n^{3/5}}} \left(\frac{R_v}{\pi_v} e^{-(1+O(T\pi_v))t\pi_v/R_v} + O(\lambda e^{-\lambda t/2}) \right).$$

Let $t_1 = (1 + \epsilon)t_0$ where $\epsilon = n^{-1/3}$ can be assumed by Lemma 10 to satisfy $T\pi_v = o(\epsilon)$ for all $v \in V$.

If $v \notin V_B$, $v \geq n^{3/5}$ then by Lemmas 16(a) and 17(a),

$$(38) \quad t_1(1 + O(T\pi_v))\pi_v/R_v \geq \frac{2m}{m-1} n \ln n \cdot \frac{d(v)}{2mn} \cdot \frac{m-1}{d(v)} = \ln n.$$

Plugging (38) into (26) and using $R_v \leq 5$ (Lemmas 16(b) and 17(b)) and $\pi_v \geq \frac{1}{2n}$ for all $v \in V \setminus V_B$ we get

$$(39) \quad C_u \leq t_1 + 10n + o(n) + \sum_{v \in V_B} \sum_{s \geq t} \Pr(\mathbf{A}_s(v))$$

It remains to deal with $v \in V_B, v \geq n^{3/5}$. We first observe that

$$(40) \quad |V_B| \leq (\ln n)^{5\omega} \omega^{3\omega} \leq (\omega \ln n)^{5\omega}$$

Using Lemma 18 we have

$$\sum_{v \in V_B} \sum_{s \geq t_1} \Pr(\mathbf{A}_s(v)) \leq (\omega \ln n)^{5\omega} \left(2n\omega e^{-(1+o(1))t_1/(2n(\ln n)^{2/3})} + O(\lambda e^{-\lambda t_1/2}) \right) = o(n).$$

This completes our proof of the upper bound on cover time.

6.3 Lower bound on cover time

Done in a similar way to that for regular graphs.

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