

# An Upper Bound for the Number of Complete Mappings

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## Abstract

A *complete mapping* on a set  $G$  with binary operation  $\circ$  is a bijection  $\theta : G \rightarrow G$  such that the mapping  $\eta : G \rightarrow G$  defined by  $\eta(x) = x \circ \theta(x)$  is again a bijection.

We give an asymptotic upper bound of  $\exp\{-0.08854n\}$  for the proportion of permutations in  $S_n$  which form complete mappings under addition modulo  $n$ .

## 1 Introduction

Following the notation of Dénes and Keedwell [2], we define a *complete mapping* on a set  $G$  with binary operation  $\circ$  as a bijection  $\theta : G \rightarrow G$  such that the mapping  $\eta : G \rightarrow G$  defined by  $\eta(x) = x \circ \theta(x)$  is again a bijection.

We restrict our attention here to the case where  $(G, \circ)$  is  $(\mathbf{Z}_n, +)$  the group of integers under addition modulo  $n$ . For example, in  $(\mathbf{Z}_5, +)$ , the mapping  $\theta$ , given below, is complete with associated mapping  $\eta(x) = (x + \theta(x)) \pmod{5}$ , given by respectively

$$\theta = \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ 0 & 1 & 2 & 3 & 4 \end{pmatrix} \quad \text{and} \quad \eta = \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ 0 & 2 & 4 & 1 & 3 \end{pmatrix}.$$

The set of bijections from  $\mathbf{Z}_n$  to  $\mathbf{Z}_n$  is  $S_n$ , the set of permutations on  $\{0, 1, \dots, n-1\}$ . We ask the following natural question:

If  $G_n = \{\sigma \in S_n : \sigma \text{ is a complete mapping on } (\mathbf{Z}_n, +)\}$ , what can we say about  $|G_n|/|S_n|$ ?

Any group  $G$  of odd order has a complete mapping (see Theorem 1.4.3 of [2]) this being a consequence of the fact that every element of  $G$  has a unique square root. In the case of

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$(\mathbf{Z}_n, +)$ ,  $n$  odd, it is trivial to construct  $n$  complete maps based on the identity permutation, by cyclically rotating the bottom row.

When  $n$  is even, there are no complete mappings on  $(\mathbf{Z}_n, +)$  as may be seen from the following theorem of I.J. Paige [4] (quoted from Theorem 1.4.5 of [2])

**Theorem 1** *If  $(G, \circ)$  is a finite group of order  $n$  which has a complete mapping, there exists an ordering of its elements, say,  $a_1, a_2, \dots, a_n$  such that  $a_1 a_2 \cdots a_n = e$ , where  $e$  is the identity of  $G$ .*

Considering  $(\mathbf{Z}_n, +)$  we can only have  $\frac{1}{2}n(n-1) \equiv 0 \pmod{n}$  if  $n$  divides  $\frac{1}{2}n(n-1)$ , which is true only when  $n$  is odd.

Our main result is the following upper bound.

**Theorem 2** *There exists a constant  $c \geq 0.08854$  such that for sufficiently large  $n$ ,*

$$\frac{|G_n|}{|S_n|} \leq \exp\{-cn\}.$$

We offer the reader two proofs of this theorem. The first uses a martingale inequality to establish a slightly weaker value of 0.06766 for  $c$ . A subsequent analysis from first principles improves the value of  $c$  to 0.08854. We are at present unable to further improve this bound, although intuitive considerations of Poisson occupancy would suggest that  $c$  is at least 1.

For the purposes of discussion, we refer to the set  $G_n$  as the set of *good* permutations.

## 2 Proof that $c$ is at least 0.06766

The proof is an application of a martingale inequality which may be found (for example) in the survey paper by McDiarmid [3]. For completeness we give the necessary definitions and statement of the inequality.

Let  $(V, d)$  be a finite metric space. A *partition sequence*  $((\mathcal{P}_k, c_k) : k = 0, \dots, n)$  consists of a sequence  $\mathcal{P}_0, \mathcal{P}_1, \dots, \mathcal{P}_n$  of *increasingly refined partitions* of  $V$ , starting with the trivial partition  $\mathcal{P}_0$  with single subset  $V$ , and ending with the discrete partition  $\mathcal{P}_n$  of  $V$  into singleton sets, and with the general property that if  $A \in \mathcal{P}_k$  then  $A \subseteq B \in \mathcal{P}_{k-1}$ .

The sequence  $c_0, c_1, \dots, c_k, \dots, c_n$  of positive reals has the following property. For each  $k$ , ( $k = 1, \dots, n$ ), whenever  $A, B \in \mathcal{P}_k$  and  $A, B \subseteq C \in \mathcal{P}_{k-1}$  for some  $C$ , then there is a bijection  $\phi : A \rightarrow B$  with  $d(x, \phi(x)) \leq c_k$  for all  $x \in A$ .

For the case we consider here,  $V = S_n$ , and if  $\alpha, \beta \in S_n$ , where  $\alpha = (\alpha(i) : i = 1, \dots, n)$  we use the metric

$$d(\alpha, \beta) = \sum_{j:\alpha(j) \neq \beta(j)} 1.$$

Two permutations  $\alpha, \beta$  are placed in the same subset  $A$  in the partition  $\mathcal{P}_k$  of  $S_n$  if they are identical on the first  $k$  entries. In other words  $\alpha(1) = \beta(1), \dots, \alpha(k) = \beta(k)$ . Clearly such a partition is increasingly refined, for if  $A, B \in \mathcal{P}_k$  where  $A \neq B$  but  $A, B \subset C \in \mathcal{P}_{k-1}$ , then the elements of  $\alpha$  and  $\beta$  differ for the first time at the  $k$ -th entry. Of course every  $\alpha \in A$  has identical value  $\alpha(k)$ , and similarly for every  $\beta \in B$ .

Any  $\alpha \in A$  is such that  $\alpha(i) = \beta(k)$  for some  $i > k$ , and we define  $\phi(\alpha)$  to be that  $\beta \in B$  for which the elements  $\alpha(k), \alpha(i)$  are transposed and all other elements remain unaltered. But  $d(\alpha, \phi(\alpha)) \leq 2$ , as  $\alpha$ , and  $\phi(\alpha)$  are identical apart from the stated transposition, and thus  $c_k = 2$ .

**Theorem 3** *Suppose that the finite metric space  $(V, d)$  has a partition sequence  $((\mathcal{P}_k, c_k) : k = 0, \dots, n)$ . Let the function  $f$  on  $V$  satisfy  $|f(x) - f(y)| \leq d(x, y)$  for all  $x, y \in V$ . Let  $X$  be uniformly distributed over  $V$ . Then for any  $t > 0$ ,*

$$\Pr(f(X) - E(f(X)) \geq t) \leq \exp\{-2t^2 / \sum c_k^2\}.$$

In our case let  $J_\alpha = \{j \in \{0, 1, \dots, n-1\} : i + \alpha(i) = j\}$ , and set  $f(\alpha) = |J_\alpha|$ . Note that  $S_n$  has the uniform (counting) measure. We must show that our choice of  $f$  satisfies

$$|f(\alpha) - f(\beta)| \leq d(\alpha, \beta), \quad (1)$$

for any choice of  $\alpha, \beta$ . Denote  $J_\alpha \cap J_\beta$  by  $J_{\alpha\beta}$  and  $J_\alpha \cap \bar{J}_\beta$  by  $J_{\alpha\bar{\beta}}$ . Let  $j \in J_{\alpha\bar{\beta}}$ . Then for some  $i$  we have  $i + \alpha(i) = j$ , and obviously  $i + \beta(i) \neq j$ . Thus  $\alpha(i) \neq \beta(i)$ . The indices  $i$  are distinct for  $j \in J_{\alpha\bar{\beta}}$ , and therefore

$$|J_{\alpha\bar{\beta}}| \leq d(\alpha, \beta).$$

On the other hand,

$$f(\alpha) - f(\beta) = |J_{\alpha\bar{\beta}}| - |J_{\bar{\alpha}\beta}| \leq |J_{\alpha\bar{\beta}}|.$$

We conclude that

$$f(\alpha) - f(\beta) \leq d(\alpha, \beta),$$

and symmetry implies the truth of (1).

The result that  $E(|J|)$  is asymptotic to  $n(1 - e^{-1})$  follows from the result that the number of  $i + \alpha(i)$  equal to  $0 \pmod{n}$  is asymptotically Poisson with parameter 1. This may be seen by considering  $k$ -th factorial moments or appealing directly to Theorem 4A of [1] which bounds the error term in this approximation by  $2/n$ .

Thus

$$\begin{aligned} \Pr(\alpha \in G_n) &= \Pr(|J_\alpha| = n) \\ &\leq \Pr\left(|J_\alpha| - E(|J|) \geq ne^{-1}(1 + o(1))\right) \\ &\leq \exp\left\{-\left(\frac{e^{-2}}{2}\right)(1 + o(1))n\right\}, \end{aligned}$$

as required.

### 3 Proof that $c$ is at least 0.08854

We first describe a method for generating permutations by sampling without replacement. Choose  $(x_1, y_1)$  at random so that

$$\Pr(x_1 = i, y_1 = j) = \frac{1}{n^2} \quad 0 \leq i, j \leq n - 1,$$

and set  $X_1 = \{x_1\}, Y_1 = \{y_1\}$ . In general, choose  $x_k$  from  $\{0, \dots, n - 1\} \setminus X_{k-1}$ , where  $X_{k-1} = \{x_1, \dots, x_{k-1}\}$ , and similarly for  $y_k$ . Thus

$$\Pr((x_k, y_k) = (i, j) \mid X_{k-1}, Y_{k-1}) = \frac{1}{(n - (k - 1))^2},$$

for  $i \notin X_{k-1}, j \notin Y_{k-1}$ . As long as no confusion arises we will identify the set  $X_k$  with the vector  $(x_1, \dots, x_k)$ .

Set  $z_k = x_k + y_k \pmod{n}$ , and say that a failure occurs at time  $k$  if  $z_k \in Z_{k-1}$ , where  $Z_{k-1} = \{z_1, \dots, z_{k-1}\}$ .

If no failure occurs at times  $1, \dots, n$  then the permutation is good. Evidently,

$$|G_n| = n! E [p_2(X_1, Y_1)p_3(X_2, Y_2)\dots p_n(X_{n-1}, Y_{n-1})],$$

where  $p_k(X_{k-1}, Y_{k-1})$  is the conditional probability of success at time  $k$  given  $(X_{k-1}, Y_{k-1})$ .

The entire set  $I = \{(i, j) : 0 \leq i, j \leq n - 1\}$  consists of *admissible* and *inadmissible* positions for the new point  $(x_k, y_k)$ . It is easy to see there are  $(n - (k - 1))^2$  admissible positions  $(i, j)$ . We now estimate how many of these admissible positions satisfy  $i + j \pmod{n} \notin Z_{k-1}$ .

For a point  $(i, j) \in I$  each of the following three events may occur,

$$\begin{aligned} A &= \{i \in X_{k-1}\} \\ B &= \{j \in Y_{k-1}\} \\ C &= \{i + j \pmod{n} \in Z_{k-1}\}. \end{aligned}$$

Denote by  $N(A)$  then number of points in the set  $A$ , and similarly for  $N(B), N(C)$  etc. We observe that

$$p_k(X_{k-1}, Y_{k-1}) = \frac{n^2 - N(A \cup B \cup C)}{(n - k + 1)^2}.$$

By the Inclusion-Exclusion principle, we have

$$N(A \cup B \cup C) = N(A) + N(B) + N(C) - N(AB) - N(BC) - N(AC) + N(ABC).$$

All the terms except the last one are easy to obtain precisely, namely

$$N(A) = N(B) = N(C) = n(k - 1),$$

$$N(AB) = N(AC) = N(BC) = (k-1)^2.$$

For the last term, one can at least assert the trivial inequality,

$$(k-1) \leq N(ABC) \leq (k-1)^2,$$

and thus

$$1 - \frac{k-1}{n-k+1} \leq p_k(X_{k-1}, Y_{k-1}) \leq 1 - \frac{(k-1)(n-2k+3)}{(n-k+1)^2}.$$

Considering the upper bound, one obtains,

$$\begin{aligned} E(p_2(X_1, Y_1) \dots p_n(X_{n-1}, Y_{n-1})) &\leq \prod_{1 \leq k \leq (n+3)/2} \left( 1 - \frac{(k-1)(n-2k+3)}{(n-k+1)^2} \right) \\ &= \exp \left\{ \sum_{1 \leq k \leq (n+3)/2} \ln \left( 1 - \frac{(k-1)(n-2k+3)}{(n-k+1)^2} \right) \right\} \\ &\leq \exp \left\{ (1+o(1))n \int_0^{1/2} \ln \left( 1 - \frac{x(1-2x)}{(1-x)^2} \right) dx \right\}. \end{aligned}$$

Simple inequalities allow the integral to be bounded by  $(3 \ln 2 - 2)$  and numerical integration gives a more precise value for the integral of  $-0.0885474$ .

## 4 Counting good permutations in a restricted class

Let  $\alpha = \{(i, \alpha(i)), i = 1, \dots, n\}$  be a permutation, and denote by  $N_k(a)$  the number of entries  $(i, \alpha(i))$  of  $\alpha$  occurring in the rectangle  $\{1, \dots, k\} \times \{a, \dots, a+m-1\}$

We define the deviation from the expected value

$$\Delta_k(a) = N_k(a) - \frac{km}{n},$$

and maximal deviation

$$\Delta_m^+ = \max \left\{ \Delta_k(a) : a \in \mathbf{Z}_n, k = m, 2m, \dots, \lfloor \frac{n-m}{m} \rfloor m \right\}.$$

Let  $G_n(z)$  be the subset of good permutations  $G_n$  for which  $\Delta_m^+ \leq z$ .

**Theorem 4** *Let  $n \rightarrow \infty$  and let  $m, z$  vary in such a way that  $m = o(n)$ ,  $z = o(m)$ . Then there exists  $\epsilon > 0$ , such that for sufficiently large  $n$ ,*

$$\frac{|G_n(z)|}{|S_n|} \leq e^{-(1-\epsilon)n}.$$

**Proof** Let  $\alpha \in G_n(z)$ . Assume that  $\alpha(1), \dots, \alpha(rm)$  are already chosen. How many possibilities exist for the choice of  $\alpha(rm+1), \dots, \alpha(rm+m)$ ? This number equals

$$mn - 2m^2r + L, \quad (2)$$

where  $L$  is the number of points  $(i, j)$  in the rectangle

$$\{rm+1, \dots, rm+m\} \times \{1, \dots, n\}$$

for which

$$j \in \{\alpha(1), \dots, \alpha(rm)\},$$

and

$$i + j \in \{\nu + \alpha(\nu) \pmod{n}, 1 \leq \nu \leq rm\}.$$

Thus  $L$  equals the number of intersections of  $rm$  ‘horizontal’ and  $rm$  ‘inclined’ segments.

There are  $N_{rm}(a)$  intersections along an inclined segment incident with the point  $(rm+m, a)$ . Thus if

$$a_\nu = [\nu + \alpha(\nu) - (rm+m)] \pmod{n},$$

then

$$L = N_{rm}(a_1) + \dots + N_{rm}(a_{rm}),$$

and

$$L \leq \frac{r^2m^3}{n} + rm\Delta_m^+.$$

Let  $x_i$  denote the number of choices for  $\alpha(rm+i)$ , then

$$\max\{x_1x_2 \cdots x_m : x_1 + \dots + x_m = c, x_1 \geq 0, \dots, x_m \geq 0\}$$

is achieved at the point  $x_i = c/m$ ,  $1 \leq i \leq m$ , where  $c$  is given by (2). Hence the number of possibilities for the choice of  $\alpha(rm+1), \dots, \alpha(rm+m)$  does not exceed

$$\begin{aligned} \left(n - 2mr + \frac{r^2m^2}{n} + r\Delta_m^+\right)^m &= n^m \left(\left(1 - \frac{rm}{n}\right)^2 + \frac{r}{n}\Delta_m^+\right)^m \\ &\leq n^m \left(\left(1 - \frac{rm}{n}\right)^2 + \frac{rz}{n}\right)^m. \end{aligned}$$

Hence for  $sm \leq n$ ,

$$|G_n(z)| \leq n^{sm} \prod_{r=0}^{s-1} \left(\left(1 - \frac{rm}{n}\right)^2 + \frac{rz}{n}\right)^m (n-sm)!, \quad (3)$$

where  $(n-sm)!$  bounds the number of choices for  $\alpha(n-sm+1), \dots, \alpha(n)$ .

Choose  $\delta > 0$ ,  $0 < \delta < 1$ , and let  $s$  be such that

$$sm \leq (1-\delta)n \leq (s+1)m.$$

Then

$$m \sum_{r=0}^{s-1} \ln \left( \left(1 - \frac{rm}{n}\right)^2 + \frac{rz}{n} \right) = n(I_0 + I_1),$$

where

$$\begin{aligned} I_0 &= \frac{2m}{n} \sum_{r=0}^{s-1} \ln \left(1 - \frac{rm}{n}\right), \\ I_1 &= \frac{m}{n} \sum_{r=0}^{s-1} \ln \left(1 + \frac{rz}{n} \left(1 - \frac{rm}{n}\right)^{-2}\right). \end{aligned}$$

Since  $m = o(n)$  we have

$$I_0 \sim 2 \int_0^{1-\delta} \ln(1-t) dt,$$

where

$$\int_0^1 \ln(1-t) dt = -1.$$

Hence  $\delta$  can be chosen in such a way that

$$I_0 < -2 + \frac{\epsilon}{2}. \quad (4)$$

For sufficiently large  $n$

$$\begin{aligned} I_1 &\leq \frac{m}{n} \sum_{r=0}^{s-1} \frac{rz}{n\delta^2} \\ &\sim \frac{zm}{2n^2\delta^2} s^2 \\ &\sim \frac{zm}{2n^2\delta^2} \left( \frac{n(1-\delta)}{m} \right)^2 \\ &< \frac{z}{m\delta^2}. \end{aligned}$$

Since  $z = o(m)$  we have

$$I_1 < \frac{\epsilon}{2}. \quad (5)$$

Setting (4) and (5) into (3) and noting that  $(n-sm)! \leq n^{n-sm}$  we obtain

$$|G_n(z)| < n^n e^{-(2-\epsilon)n},$$

for  $n$  sufficiently large. Hence for  $n$  sufficiently large,

$$\frac{|G_n(z)|}{|S_n|} < e^{-(1-\epsilon)n},$$

and the theorem follows. □

We now consider the following question.

Let  $\alpha$  be a random permutation. How can the probability of the event  $\{\Delta_m^+ \leq z\}$  be estimated?

For the complement of the event we have

$$\Pr(\Delta_m^+ > z) \leq n^s \max_{0 \leq k \leq n} \left\{ \Pr(N_k(0) > \frac{km}{n} + z) \right\}.$$

Let  $p(i) = \Pr(N_k(0) = i)$ . From a combinatorial argument we have

$$p(i) = \binom{m}{i} \frac{(k)_i (n-k)_{m-i}}{(n)_m}.$$

Setting  $c = n/(n-m)$ ,  $k = nt$ ,  $\tau = 1-t$ , we obtain

$$p(i) \leq \binom{m}{i} t^i (\tau c)^{m-i},$$

and

$$\sum_{i=0}^m e^{\lambda i} p(i) \leq (te^\lambda + \tau c)^m,$$

for arbitrary real  $\lambda$ . Hence if  $\lambda > 0$ ,

$$\Pr(N_k(0) \geq u) \leq e^{-\lambda u} (te^\lambda + \tau c)^m.$$

Setting  $u = mt + z$ ,  $e^\lambda = c(mt\tau + z\tau)/(mt\tau - zt)$ , and assuming that  $z = o(m)$ , we obtain

$$\ln \Pr(N_k(0) \geq tm + z) \leq m \left( \frac{\tau m}{n-m} - \frac{z^2}{2m^2 + \tau} \right) (1 + o(1)),$$

and hence

$$\ln \Pr(\Delta_m^+ \geq z) \leq \left( \frac{n}{m} \ln n + \frac{m^2}{n-m} - \frac{2z^2}{m} \right) (1 + o(1)).$$

For example, assume that  $m \sim n\sigma$ ,  $z \sim m\sqrt{\sigma}$  when  $\sigma = \sigma_m$  is a slowly decreasing function, then

$$\Pr(\Delta_m^+ \geq z) \leq \exp \left\{ -\sigma^2 n (1 + o(1)) \right\}.$$

Hence if  $z$  is chosen appropriately, a random permutation which satisfies  $\Delta_m^+ \leq z$  belongs to  $G_n(z)$  with probability tending to 1.

## 5 References

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