

Protein interaction networks via Tailored graph ensembles: a study of sampling

A Annibale

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Outline

- 1 Motivation
- 2 Quantifying biases
 - Tailored random graph ensembles
 - Sampling protocols
 - Results
- 3 Inferring the true network from imperfect data
 - Bayesian analysis

Motivation

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Commentary

Nature Biotechnology **26**, 69 - 72 (2008)

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Protein-protein interaction networks and biology—what's the connection?

Luke Hakes¹, John W Pinney¹, David L Robertson¹ & Simon C Lovell¹

Analysis of protein-protein interaction networks is an increasingly popular means to infer biological insight, but is close enough attention being paid to data handling protocols and the degree of bias in the data?

The availability of large-scale protein-protein interaction data has led to the recent popularity of the study of protein interaction networks. Just as the enormous amount of available sequence data has made it

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- Large-scale PIN dataset are available nowadays
- Data are biased and incomplete

Is it possible to use the available data reliably?

Yes, if we understand the **relation** between the patterns of a real graph and those of a graph sample

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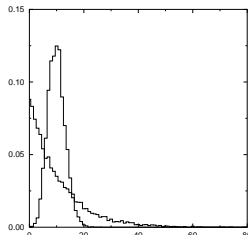
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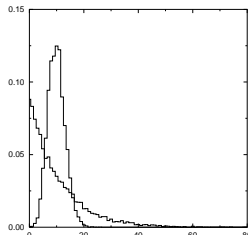
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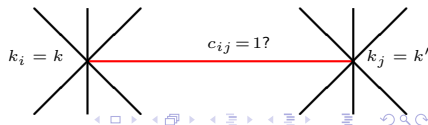
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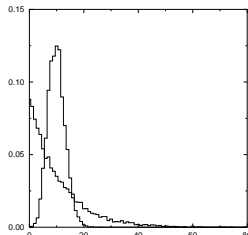
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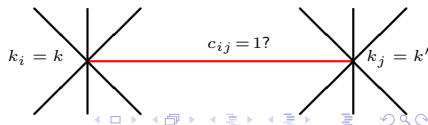
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$\Pi \neq 1$ signals presence of structure beyond degrees statistics

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Numerical sampling of $\omega_\mu(\Omega)$ hard for sophisticated Ω , but **analytical** progress feasible for suitable choices and N large!

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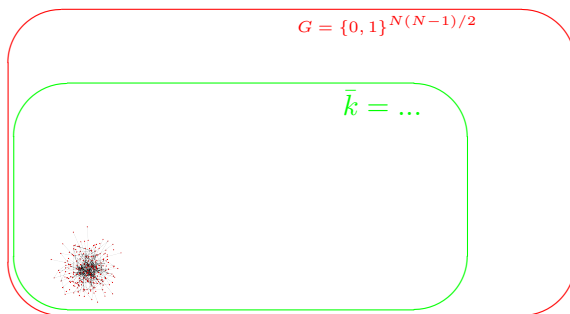
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$$G = \{0, 1\}^{N(N-1)/2}$$



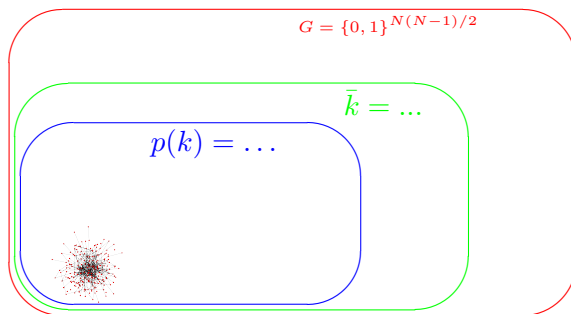
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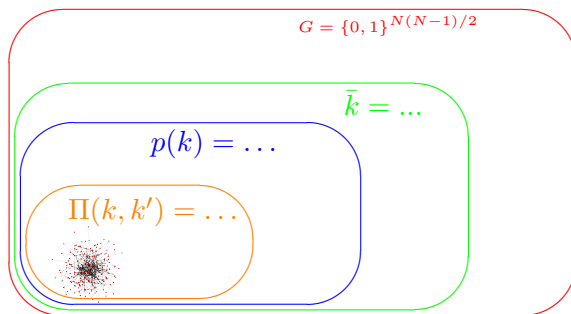
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(i) prescribe only $\langle k \rangle$ Erdős-Renyi graphs

$$P(\mathbf{c}|\bar{k}) = \prod_{i < j} \left[\frac{\bar{k}}{N} \delta_{c_{ij},1} + \left(1 - \frac{\bar{k}}{N} \right) \delta_{c_{ij},0} \right] \quad (\text{soft})$$

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$$P(\mathbf{c}|\mathbf{k}, Q) = \frac{\delta_{\mathbf{k},\mathbf{k}(\mathbf{c})}}{Z(\mathbf{k}, Q)} \prod_{i < j} \left[\frac{\langle k \rangle}{N} Q(k_i, k_j) \delta_{c_{ij},1} + \left(1 - \frac{\langle k \rangle}{N} Q(k_i, k_j) \right) \delta_{c_{ij},0} \right]$$

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$$Q(k, k') = \Pi(k, k') k k' / \langle k \rangle^2 = W(k, k') / p(k) p(k')$$

Modelling biological networks

- Biological network **c**

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[A Annibale, ACC Coolen, LP Fernandes, F Fraternali J Kleinjung *J. Phys. A: Math. Theor.* 42 485001 (2009)]

Distance between networks

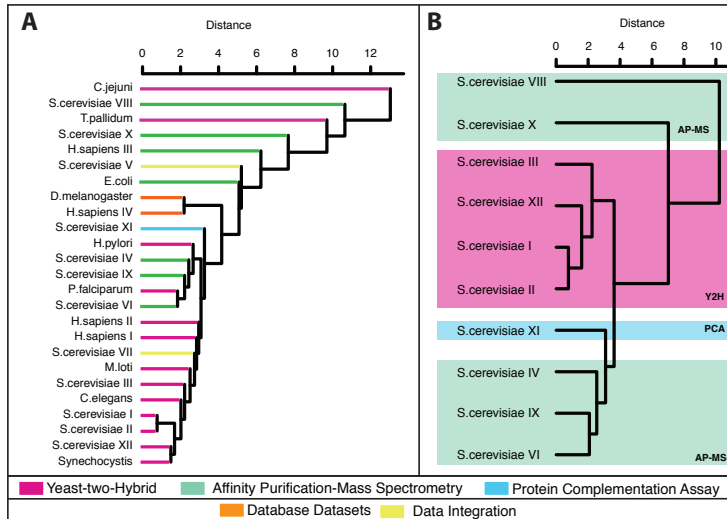
Information theory



Distance between \mathbf{c}_A and \mathbf{c}_B

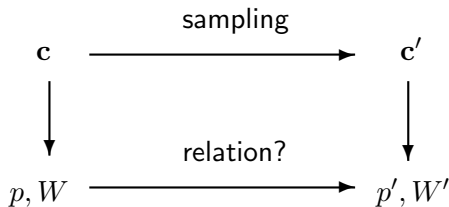
$$\begin{aligned} D_{AB} &= \frac{1}{2N} \sum_{\mathbf{c}} P(\mathbf{c}|p_A, W_A) \log \frac{P(\mathbf{c}|p_A, W_A)}{P(\mathbf{c}|p_B, W_B)} \\ &\quad + \frac{1}{2N} \sum_{\mathbf{c}} P(\mathbf{c}|p_B, W_B) \log \frac{P(\mathbf{c}|p_B, W_B)}{P(\mathbf{c}|p_A, W_A)} \\ &= f(p_A, p_B, W_A, W_B) \end{aligned}$$

Dendrograms

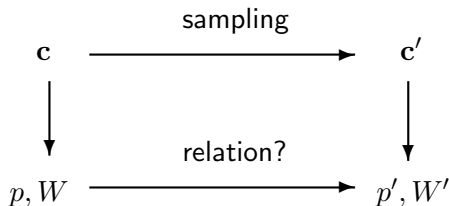


[LP Fernandes, A Annibale, J Kleinjung, ACC Coolen, F Fraternali *PLoS ONE* 5(8): e12083 (2010)]

Accounting for biases



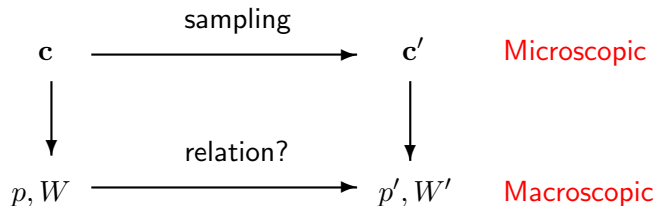
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So far:

- only p' was studied and
- only for random node sampling

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$$\tau_{ij} = \begin{cases} 1 & \text{bond } i - j \text{ detected} \\ 0 & \text{otherwise} \end{cases} \quad \begin{matrix} y \\ 1 - y \end{matrix}$$

- bond oversampling: $c'_{ij} = c_{ij} + (1 - c_{ij}) \lambda_{ij}$

$$\lambda_{ij} = \begin{cases} 1 & \text{bond } i - j \text{ created} \\ 0 & \text{otherwise} \end{cases} \quad \begin{matrix} N^{-1}z \\ 1 - N^{-1}z \end{matrix}$$

\Downarrow in combination

$$c'_{ij} = \sigma_i \sigma_j [\tau_{ij} c_{ij} + (1 - c_{ij}) \lambda_{ij}]$$

(Connectivity-dependent) Sampling protocols

- node undersampling: $c'_{ij} = \sigma_i \sigma_j c_{ij}$

$$\sigma_i = \begin{cases} 1 & \text{node } i \text{ detected} \\ 0 & \text{otherwise} \end{cases} \quad \begin{matrix} x(k_i) \\ 1 - x(k_i) \end{matrix}$$

- bond undersampling: $c'_{ij} = \tau_{ij} c_{ij}$

$$\tau_{ij} = \begin{cases} 1 & \text{bond } i - j \text{ detected} \\ 0 & \text{otherwise} \end{cases} \quad \begin{matrix} y(k_i, k_j) \\ 1 - y(k_i, k_j) \end{matrix}$$

- bond oversampling: $c'_{ij} = c_{ij} + (1 - c_{ij}) \lambda_{ij}$

$$\lambda_{ij} = \begin{cases} 1 & \text{bond } i - j \text{ created} \\ 0 & \text{otherwise} \end{cases} \quad \begin{matrix} N^{-1} z(k_i, k_j) \\ 1 - N^{-1} z(k_i, k_j) \end{matrix}$$

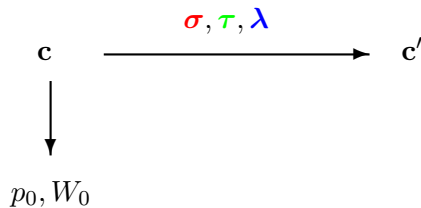
\Downarrow in combination

$$c'_{ij} = \sigma_i \sigma_j [\tau_{ij} c_{ij} + (1 - c_{ij}) \lambda_{ij}]$$

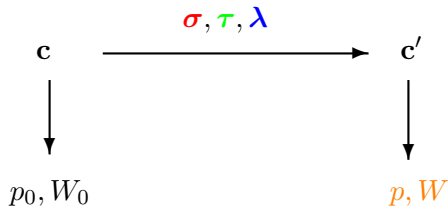
Macroscopic features

$$\mathbf{c} \xrightarrow{\sigma, \tau, \lambda} \mathbf{c}'$$

Macroscopic features



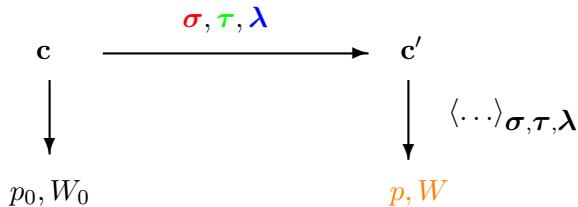
Macroscopic features



$$p(k|\mathbf{c}') = \frac{\sum_i \sigma_i \delta_{k, \sum_j c_{ij}'}}{\sum_i \sigma_i}$$

$$W(k, k'|\mathbf{c}') = \frac{\sum_{ij} c'_{ij} \delta_{k, \sum_\ell c'_{i\ell}} \delta_{k', \sum_\ell c'_{j\ell}}}{\sum_{ij} c'_{ij}}$$

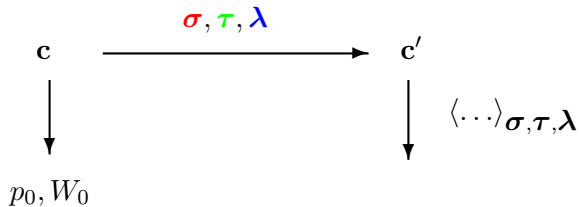
Macroscopic features



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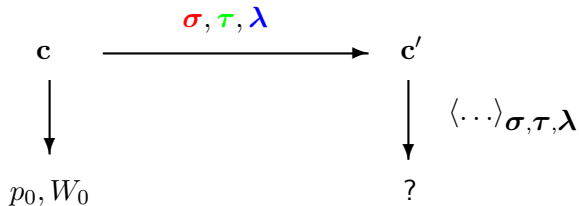
Macroscopic features



$$p(k|\mathbf{c}') = \left\langle \frac{\sum_i \sigma_i \delta_{k, \sum_j c_{ij}'} }{\sum_i \sigma_i} \right\rangle_{\sigma, \tau, \lambda}$$

$$W(k, k'|\mathbf{c}') = \left\langle \frac{\sum_{ij} c'_{ij} \delta_{k, \sum_{\ell} c'_{i\ell}} \delta_{k', \sum_{\ell} c'_{j\ell}} }{\sum_{ij} c'_{ij}} \right\rangle_{\sigma, \tau, \lambda}$$

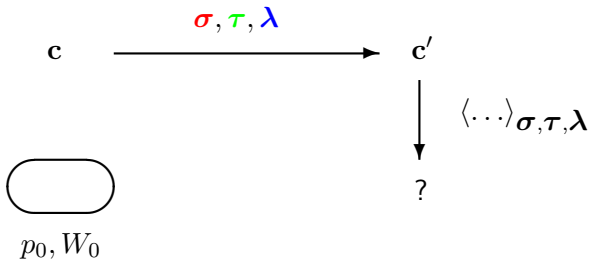
Macroscopic features



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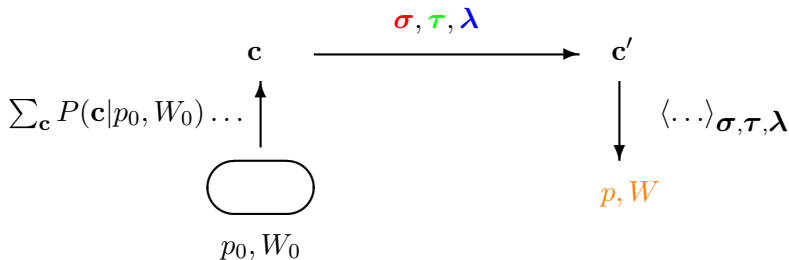
Macroscopic features



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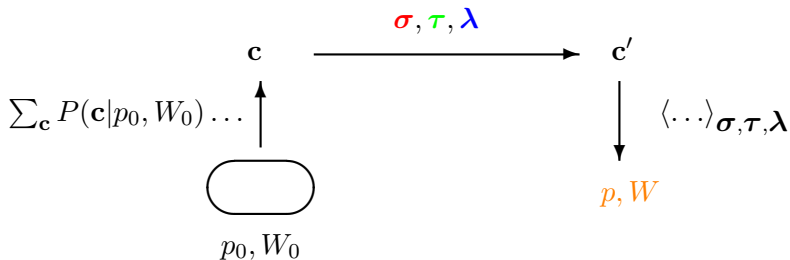
Macroscopic features



$$p(k|\mathbf{x}, \mathbf{y}, \mathbf{z}) = \lim_{N \rightarrow \infty} \sum_{\mathbf{c}} P(\mathbf{c}|p_0, W_0) \left\langle \frac{\sum_i \sigma_i \delta_{k, \sum_j c_{ij}'}}{\sum_i \sigma_i} \right\rangle_{\sigma, \tau, \lambda}$$

$$W(k, k'|\mathbf{x}, \mathbf{y}, \mathbf{z}) = \lim_{N \rightarrow \infty} \sum_{\mathbf{c}} P(\mathbf{c}|p_0, W_0) \left\langle \frac{\sum_{ij} c'_{ij} \delta_{k, \sum_{\ell} c'_{i\ell}} \delta_{k', \sum_{\ell} c'_{j\ell}}}{\sum_{ij} c'_{ij}} \right\rangle_{\sigma, \tau, \lambda}$$

Macroscopic features

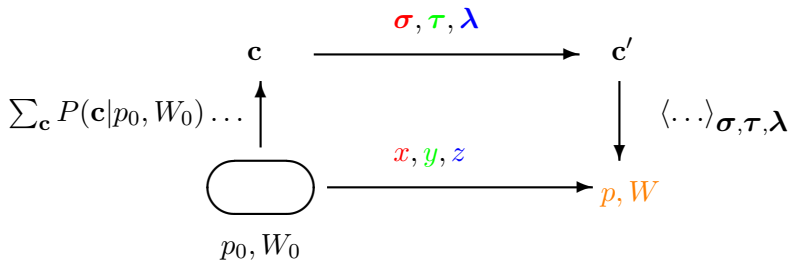


$$p(k|\mathbf{x}, \mathbf{y}, \mathbf{z}) = \lim_{N \rightarrow \infty} \sum_{\mathbf{c}} P(\mathbf{c}|p_0, W_0) \left\langle \frac{\sum_i \sigma_i \delta_{k, \sum_j c_{ij'}}}{\sum_i \sigma_i} \right\rangle_{\sigma, \tau, \lambda}$$

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\Rightarrow Statistical mechanics techniques \Rightarrow

Macroscopic features



$$p(k|x, y, z) = \lim_{N \rightarrow \infty} \sum_{\mathbf{c}} P(\mathbf{c}|p_0, W_0) \left\langle \frac{\sum_i \sigma_i \delta_{k, \sum_j c_{ij}}}{\sum_i \sigma_i} \right\rangle_{\sigma, \tau, \lambda}$$

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\Rightarrow Statistical mechanics techniques \Rightarrow

Outline

1 Motivation

2 Quantifying biases

- Tailored random graph ensembles
- Sampling protocols
- Results

3 Inferring the true network from imperfect data

- Bayesian analysis

Results

$$p(k|x, y, z) = \frac{\sum_q x(q)p(q) \{a(q)\mathcal{J}(k|q) + qb(q)\mathcal{L}(k|q)\}}{k \sum_q p(q)x(q)}$$

$$W(k, k'|x, y, z) = \frac{\sum_{q, q' > 0} x(q)x(q') \{p(q)p(q')z(q, q')\mathcal{J}(k|q)\mathcal{J}(k'|q') + \bar{k}W(q, q')y(q, q')\mathcal{L}(k|q)\mathcal{L}(k'|q')\}}{\bar{k}(x, y, z) \sum_q p(q)x(q)}$$

$$\bar{k}(x, y, z) = \sum_k k p(k|x, y, z) = \frac{\sum_q x(q)p(q)[a(q) + qb(q)]}{\sum_q p(q)x(q)}$$

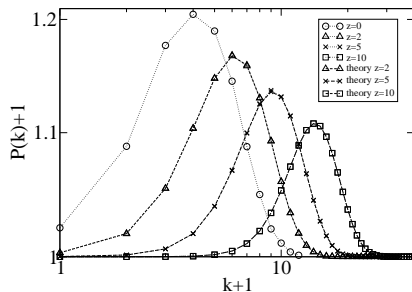
with

$$\mathcal{J}(k|q) = e^{-a(q)} \sum_{n=0}^{\min\{k-1, q\}} \binom{q}{n} \frac{a^{k-1-n}(q)}{(k-1-n)!} b^n(q) (1-b(q))^{q-n}$$

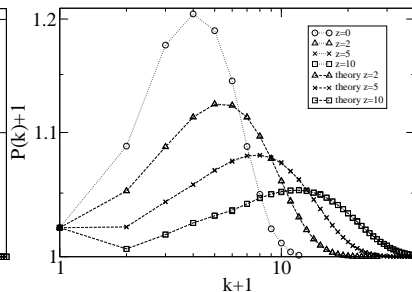
$$\mathcal{L}(k|q) = e^{-a(q)} \sum_{n=0}^{\min\{k-1, q-1\}} \binom{q-1}{n} \frac{a^{k-1-n}(q)}{(k-1-n)!} b^n(q) (1-b(q))^{q-1-n}$$

$$a(q) = \sum_{q' \geq 0} p(q')x(q')z(q, q'), \quad b(q) = \frac{\bar{k}}{qp(q)} \sum_{q' \geq 0} x(q')y(q, q')W(q, q')$$

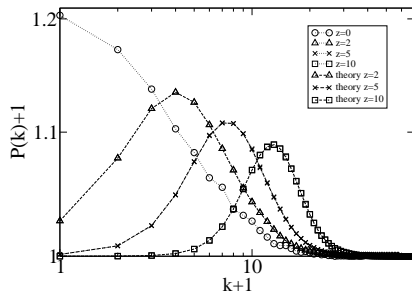
Poisson - Unbiased Bond Oversampling



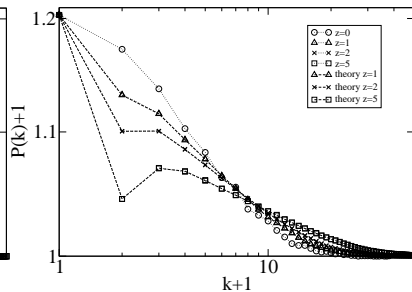
Poisson - Biased Bond oversampling



Power Law - Unbiased Bond Oversampling



Power Law - Biased Bond Oversampling



Random sampling from Elegans: degree correlations

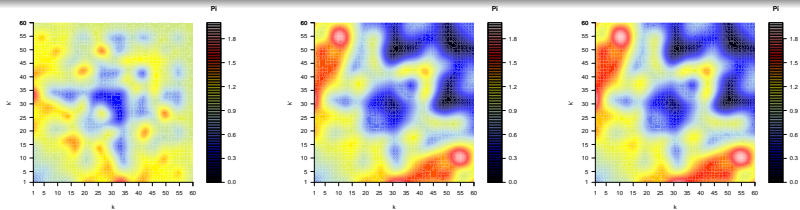


Figure: Random bond undersampling $x = 1, y = 0.9, z = 0, N = 3512, \bar{k} = 3.72$

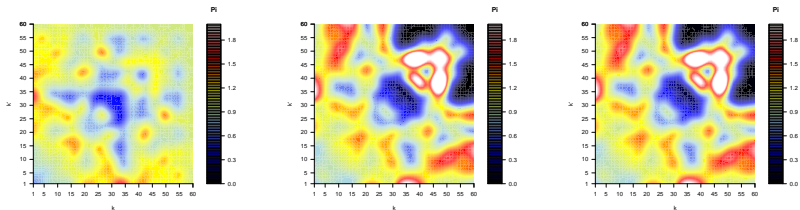


Figure: Random bond oversampling $x = 1, y = 1, z = 1, N = 3512, \bar{k} = 3.72$

[A Annibale, ACC Coolen *Interface Focus* December 6, 2011 1:836-856]

Outline

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- 2 Quantifying biases
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 - Bayesian analysis

Bayesian analysis

- $\ell = 1, \dots, L$ species

Bayesian analysis

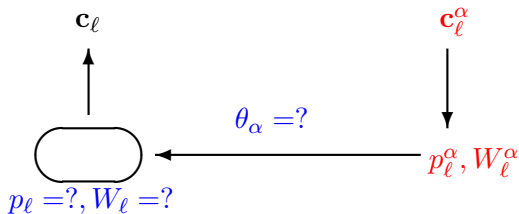
- $\ell = 1, \dots, L$ species
- $\alpha = 1, \dots, M$ experimental protocols, parameters
 $\theta_\alpha = \{x_\alpha, y_\alpha, z_\alpha\}$

Bayesian analysis

- $\ell = 1, \dots, L$ species
- $\alpha = 1, \dots, M$ experimental protocols, parameters
 $\theta_\alpha = \{x_\alpha, y_\alpha, z_\alpha\}$
- Observed networks $c_\ell^\alpha \Rightarrow p_\ell, W_\ell, \theta_\alpha?$

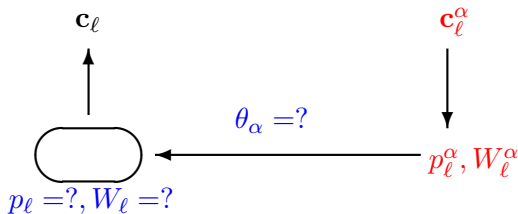
Bayesian analysis

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Bayesian analysis

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- $\alpha = 1, \dots, M$ experimental protocols, parameters
 $\theta_\alpha = \{x_\alpha, y_\alpha, z_\alpha\}$
- Observed networks $\mathbf{c}_\ell^\alpha \Rightarrow p_\ell, W_\ell, \theta_\alpha?$



- Maximize $p(\{\theta_\alpha\}, \{p_\ell\}, \{W_\ell\} | \{\mathbf{c}_\ell^\alpha\})$ over $p_\ell, W_\ell, \theta_\alpha$

- Uniform choice for prior

$$p(p_\ell, W_\ell), \quad p(\theta_\alpha)$$

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$$p(p_\ell, W_\ell), \quad p(\theta_\alpha)$$

- Likelihood:

$$p(\mathbf{c}_\ell^\alpha | \theta_\alpha, p_\ell, W_\ell) = \sum_{\mathbf{c}_\ell} P(\mathbf{c}_\ell | W_\ell, p_\ell) p(\mathbf{c}_\ell^\alpha | \theta_\alpha, \mathbf{c}_\ell)$$

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with

$p(\mathbf{c}_\ell^\alpha | \theta^\alpha, \mathbf{c}_\ell)$ determined from relation between \mathbf{c}_ℓ and \mathbf{c}_ℓ^α (known!)

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- Calculate $\langle p(\mathbf{c}_\ell^\alpha | \theta^\alpha, \mathbf{c}_\ell) \rangle$ via statistical mechanics
- Maximise posterior using Lagrange multipliers to handle constraints

$$\sum_k p_\ell(k) = 1 \qquad \sum_q W_\ell(k, q) = k p_\ell(k) / \bar{k}_\ell$$

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$$\sum_k p_\ell(k) = 1 \quad \sum_q W_\ell(k, q) = k p_\ell(k) / \bar{k}_\ell$$

- get a set of equations for p_ℓ, W_ℓ and $x^\alpha, y^\alpha, z^\alpha$ in terms of the observed $p_\ell^\alpha, W_\ell^\alpha$

Conclusions

- Tailored random graphs ensemble can be used to model complex networks and quantify distances between them.
- Tailored graph ensembles can be used to quantify sampling effects on degree distributions and degree correlations for general sampling protocols (simulations match theory!)
- Underway: Bayesian inference of macroscopic features of biological networks and sampling parameters of different experiments given the observed networks
- Future: go all the way back to the original matrices

Aknowledgements

- ACC Coolen (maths)
- LP Fernandes, F Fraternali, J Kleinjung (bioinformatics)



ACC Coolen, F Fraternali, A Annibale, LP Fernandes, J Kleinjung

Modelling Biological Networks via Tailored Random Graphs

in *Handbook of Statistical Systems Biology*, MPH Stumpf, DJ Balding, M Girolami, Wiley, 2011



A Annibale, ACC Coolen, LP Fernandes, F Fraternali J Kleinjung

Tailored graph ensembles as proxies or null models for real networks I: tools for quantifying structure

J. Phys. A: Math. Theor. **42** 485001 (2009)



ACC Coolen, A De Martino, A Annibale

Constrained Markovian dynamics of random graphs

J. Stat. Phys. **136** (2009), 1035-1067



LP Fernandes, A Annibale, J Kleinjung, ACC Coolen, F Fraternali

Protein networks reveal detection bias and species consistency when viewed through information-theoretic glasses

PLoS ONE **5**(8): e12083 (2010)



A Annibale, ACC Coolen

What you see is not what you get: how sampling affects macroscopic features of biological networks

Interface Focus (2011) **1**, 836-856



A. Annibale, ACC Coolen

Inferring protein interactions networks from biased experimental data (in preparation).